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## ► To cite this version:

Minmin Wang. Height and diameter of Brownian tree. Electronic Communications in Probability, 2015, 20, pp.1-15. 10.1214/ECP.v20-4193 . hal-01132277v2

**HAL Id: hal-01132277**

**<https://hal.sorbonne-universite.fr/hal-01132277v2>**

Submitted on 27 Jan 2016

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# Height and diameter of Brownian tree

Minmin Wang\*

## Abstract

By computations on generating functions, Szekeres proved in 1983 that the law of the diameter of a uniformly distributed rooted labelled tree with  $n$  vertices, rescaled by a factor  $n^{-1/2}$ , converges to a distribution whose density is explicit. Aldous observed in 1991 that this limiting distribution is the law of the diameter of the Brownian tree. In our article, we provide a computation of this law which is directly based on the normalized Brownian excursion. Moreover, we provide an explicit formula for the joint law of the height and diameter of the Brownian tree, which is a new result.

**Keywords:** Brownian tree; Brownian excursion; continuum random tree; Jacobi theta function; Williams' decomposition.

**AMS MSC 2010:** 60J80.

Submitted to ECP on March 18, 2015, final version accepted on November 18, 2015.

Supersedes arXiv:1503.05014v1.

## 1 Introduction

For any integer  $n \geq 1$ , let  $T_n$  be a uniformly distributed random rooted labelled tree with  $n$  vertices and we denote by  $D_n$  its diameter with respect to the graph distance. By computations on generating functions, Szekeres [22] proved that

$$n^{-\frac{1}{2}} D_n \xrightarrow{\text{(law)}} \Delta, \quad (1.1)$$

where  $\Delta$  is a random variable whose probability density  $f_\Delta$  is given by

$$f_\Delta(y) = \frac{\sqrt{2\pi}}{3} \sum_{n \geq 1} \left( \frac{64}{y^4} (4b_{n,y}^4 - 36b_{n,y}^3 + 75b_{n,y}^2 - 30b_{n,y}) + \frac{16}{y^2} (2b_{n,y}^3 - 5b_{n,y}^2) \right) e^{-b_{n,y}}, \quad (1.2)$$

where  $b_{n,y} := 8(\pi n/y)^2$ , for all  $y \in (0, \infty)$  and for all integers  $n \geq 1$ . This result is implicitly written in Szekeres [22] p. 395 formula (12). See also Broutin and Flajolet [9] for a similar result for binary trees. On the other hand, Aldous [2, 4] has proved that  $T_n$ , whose graph distance is rescaled by a factor  $n^{-\frac{1}{2}}$ , converges in distribution to the Brownian tree (also called Continuum Random Tree) that is a random compact metric space. From this, Aldous has deduced that  $\Delta$  has the same distribution as the diameter of the Brownian tree: see [3], Section 3.4, (though formula (41) there is not accurate). As proved by Aldous [4] and by Le Gall [18], the Brownian tree is coded by the normalized Brownian excursion of length 1 (see below for more detail). Then, the question was raised by Aldous [3] that whether we can establish (1.2) directly from computations

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on the normalized Brownian excursion. (See also [20], Exercise 9.4.1.) In this work, we present a solution to this question: we compute the Laplace transform for the law of the diameter of the Brownian tree based on Williams' decomposition of Brownian excursions. We also provide a formula for the joint law of the total height and diameter of the Brownian tree, which appears to be new. In a joint work with Duquesne [13], we generalize the present method to a Lévy tree and study its total height and diameter. Before stating precisely our results, let us first recall the definition of the Brownian tree coded by the normalized Brownian excursion.

**Normalized Brownian excursion.** Let  $X = (X_t)_{t \geq 0}$  be a continuous process defined on a probability space  $(\Omega, \mathcal{F}, \mathbf{P})$  such that  $(\frac{1}{\sqrt{2}} X_t)_{t \geq 0}$  is distributed as a linear standard Brownian motion such that  $\mathbf{P}(X_0 = 0) = 1$  (the reason for the normalizing constant  $\sqrt{2}$  is explained below). Thus,

$$\forall u \in \mathbb{R}, t \in \mathbb{R}_+, \quad \mathbf{E}[e^{iuX_t}] = e^{-tu^2}.$$

For all  $t \in \mathbb{R}_+$ , we set  $I_t = \inf_{s \in [0, t]} X_s$ . Then, the reflected process  $X - I$  is a strong Markov process, the state 0 is instantaneous in  $(0, \infty)$  and recurrent, and  $-I$  is a local time at level 0 for  $X - I$  (see Bertoin [6], Chapter VI). We denote by  $\mathbf{N}$  the *excursion measure* associated with the local time  $-I$ ;  $\mathbf{N}$  is a sigma finite measure on the space of continuous paths  $\mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$ . More precisely, let  $\bigcup_{i \in \mathcal{I}} (a_i, b_i) = \{t > 0 : X_t - I_t > 0\}$  be the excursion intervals of the reflected process  $X - I$  above 0; for all  $i \in \mathcal{I}$ , we set  $e_i(s) = X_{(a_i + s) \wedge b_i} - I_{a_i}$ ,  $s \in \mathbb{R}_+$ . Then,

$$\sum_{i \in \mathcal{I}} \delta_{(-I_{a_i}, e_i)} \text{ is a Poisson point measure on } \mathbb{R}_+ \times \mathbf{C}(\mathbb{R}_+, \mathbb{R}_+) \text{ with intensity } dt \mathbf{N}(de). \quad (1.3)$$

We shall denote by  $e = (e_t)_{t \geq 0}$  the canonical process on  $\mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$ . We define its *lifetime* by

$$\zeta = \sup\{t \geq 0 : e_t > 0\}, \quad (1.4)$$

with the convention that  $\sup \emptyset = 0$ . Then,  $\mathbf{N}$ -a.e.  $e_0 = 0$ ,  $\zeta \in (0, \infty)$  and for all  $t \in (0, \zeta)$ ,  $e_t > 0$ . Moreover, one has

$$\forall \lambda \in (0, \infty), \quad \mathbf{N}(1 - e^{-\lambda \zeta}) = \sqrt{\lambda} \quad \text{and} \quad \mathbf{N}(\zeta \in dr) = \frac{dr}{2\sqrt{\pi} r^{3/2}}. \quad (1.5)$$

See Blumenthal [7] IV.1 for more detail.

Let us briefly recall the scaling property of  $e$  under  $\mathbf{N}$ . To that end, recall that  $X$  satisfies the following scaling property: for all  $r \in (0, \infty)$ ,  $(r^{-\frac{1}{2}} X_{rt})_{t \geq 0}$  has the same law as  $X$ , which entails that

$$(r^{-\frac{1}{2}} e_{rt})_{t \geq 0} \text{ under } r^{\frac{1}{2}} \mathbf{N} \stackrel{(\text{law})}{=} e \text{ under } \mathbf{N}. \quad (1.6)$$

This scaling property implies that there exists a family of laws on  $\mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$  denoted by  $\mathbf{N}(\cdot | \zeta = r)$ ,  $r \in (0, \infty)$ , such that  $r \mapsto \mathbf{N}(\cdot | \zeta = r)$  is weakly continuous on  $\mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$ , such that  $\mathbf{N}(\cdot | \zeta = r)$ -a.s.  $\zeta = r$  and such that

$$\mathbf{N} = \int_0^\infty \mathbf{N}(\cdot | \zeta = r) \mathbf{N}(\zeta \in dr). \quad (1.7)$$

Moreover, by (1.6),  $(r^{-\frac{1}{2}} e_{rt})_{t \geq 0}$  under  $\mathbf{N}(\cdot | \zeta = r)$  has the same law as  $e$  under  $\mathbf{N}(\cdot | \zeta = 1)$ . To simplify notation we set

$$\mathbf{N}_{\text{nr}} := \mathbf{N}(\cdot | \zeta = 1). \quad (1.8)$$

Thus, for all measurable functions  $F: \mathbf{C}(\mathbb{R}_+, \mathbb{R}_+) \rightarrow \mathbb{R}_+$ ,

$$\mathbf{N}[F(e)] = \frac{1}{2\sqrt{\pi}} \int_0^\infty dr r^{-\frac{3}{2}} \mathbf{N}_{\text{nr}} \left[ F \left( (r^{\frac{1}{2}} e_{t/r})_{t \geq 0} \right) \right]. \quad (1.9)$$

**Remark 1.1.** The standard Ito measure  $\mathbf{N}_{\text{ito}}^+$  of positive excursions, as defined for instance in Revuz & Yor [21] Chapter XII Theorem 4.2, is derived from  $\mathbf{N}$  by the following scaling relations:

$\mathbf{N}_{\text{ito}}^+$  is the law of  $\frac{1}{\sqrt{2}}e$  under  $\frac{1}{\sqrt{2}}\mathbf{N}$  and thus,  $\mathbf{N}_{\text{ito}}^+(\cdot | \zeta=1)$  is the law of  $\frac{1}{\sqrt{2}}e$  under  $\mathbf{N}_{\text{nr}}$ .

Consequently, the law  $\mathbf{N}_{\text{nr}}$  is *not* the standard version for normalized Brownian excursion measure. However, we shall refer to it as the *normalized Brownian excursion measure*.

**Real trees.** Let us recall the definition of *real trees* that are metric spaces generalizing graph-trees: let  $(T, d)$  be a metric space; it is a real tree if the following statements hold true.

- (a) For all  $\sigma_1, \sigma_2 \in T$ , there is a unique isometry  $f: [0, d(\sigma_1, \sigma_2)] \rightarrow T$  such that  $f(0) = \sigma_1$  and  $f(d(\sigma_1, \sigma_2)) = \sigma_2$ . In this case, we set  $\llbracket \sigma_1, \sigma_2 \rrbracket := f([0, d(\sigma_1, \sigma_2)])$ .
- (b) For any continuous injective function  $q: [0, 1] \rightarrow T$ ,  $q([0, 1]) = \llbracket q(0), q(1) \rrbracket$ .

When a point  $\rho \in T$  is distinguished,  $(T, d, \rho)$  is said to be a *rooted* real tree,  $\rho$  being the *root* of  $T$ . Among connected metric spaces, real trees are characterized by the so-called *four-point inequality*: we refer to Evans [15] or to Dress, Moulton & Terhalle [11] for a detailed account on this property. Let us briefly mention that the set of (pointed) isometry classes of compact rooted real trees can be equipped with the (pointed) Gromov-Hausdorff distance which makes it into a Polish space: see Evans, Pitman & Winter [16], Theorem 2, for more detail on this intrinsic point of view that we do not adopt here.

**Coding of real trees.** Real trees can be constructed through continuous functions. Recall that  $e$  stands for the canonical process on  $\mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$ . We assume here that  $e$  has a compact support, that  $e_0 = 0$  and that  $e$  is not identically null. Recall from (1.4) the definition of its lifetime  $\zeta$ . Then, our assumptions on  $e$  entail that  $\zeta \in (0, \infty)$ . For  $s, t \in [0, \zeta]$ , we set

$$b(s, t) := \inf_{r \in [s \wedge t, s \vee t]} e_r \quad \text{and} \quad d(s, t) := e_t + e_s - 2b(s, t).$$

It turns out that  $d$  is a pseudo-distance on  $[0, \zeta]$ . We define the equivalence relation  $\sim$  by setting  $s \sim t$  iff  $d(s, t) = 0$ ; then we set

$$\mathcal{T} := [0, \zeta] / \sim. \quad (1.10)$$

The function  $d$  induces a distance on the quotient set  $\mathcal{T}$  that we keep denoting by  $d$  for simplicity. We denote by  $p: [0, \zeta] \rightarrow \mathcal{T}$  the canonical projection. Clearly  $p$  is continuous, which implies that  $(\mathcal{T}, d)$  is a compact metric space. Moreover, it is shown that  $(\mathcal{T}, d)$  is a real tree (see Duquesne & Le Gall [12], Theorem 2.1, for a proof). We take  $\rho = p(0)$  as the *root* of  $\mathcal{T}$ . The *total height* and the *diameter* of  $\mathcal{T}$  are thus given by

$$\Gamma = \max_{\sigma \in \mathcal{T}} d(\rho, \sigma) = \max_{t \geq 0} e_t \quad \text{and} \quad D = \max_{\sigma, \sigma' \in \mathcal{T}} d(\sigma, \sigma') = \max_{s, t \geq 0} (e_t + e_s - 2b(s, t)). \quad (1.11)$$

We also define on  $\mathcal{T}$  a finite measure  $m$ , called the *mass measure*, that is the pushforward measure of the Lebesgue measure on  $[0, \zeta]$  by the canonical projection  $p$ . Namely, for all continuous functions  $f: \mathcal{T} \rightarrow \mathbb{R}_+$ ,

$$\int_{\mathcal{T}} f(\sigma) m(d\sigma) = \int_0^\zeta f(p(t)) dt. \quad (1.12)$$

Note that  $m(\mathcal{T}) = \zeta$ .

**Brownian tree.** The random rooted compact real tree  $(\mathcal{T}, d, \rho)$  coded by  $e$  under the normalized Brownian excursion measure  $\mathbf{N}_{\text{nr}}$  defined in (1.8) is the *Brownian tree*. Here, we recall some properties of the Brownian tree. To that end, for any  $\sigma \in \mathcal{T}$ , we denote by  $n(\sigma)$  the number of connected components of the open set  $\mathcal{T} \setminus \{\sigma\}$ . Note that  $n(\sigma)$  is possibly infinite. We call this number the *degree* of  $\sigma$ . We say that  $\sigma$  is a *branch point* if  $n(\sigma) \geq 3$  and that  $\sigma$  is a *leaf* if  $n(\sigma) = 1$ . We denote by  $\text{Lf}(\mathcal{T}) := \{\sigma \in \mathcal{T} : n(\sigma) = 1\}$  the *set of leaves* of  $\mathcal{T}$ . Then the following holds true:

$$\mathbf{N}_{\text{nr}}\text{-a.s. } \forall \sigma \in \mathcal{T}, \quad n(\sigma) \in \{1, 2, 3\}, \quad \text{m is diffuse} \quad \text{and} \quad \text{m}(\mathcal{T} \setminus \text{Lf}(\mathcal{T})) = 0, \quad (1.13)$$

where we recall from (1.12) that  $\text{m}$  stands for the mass measure. The Brownian tree has therefore only binary branch points (i.e. branch points of degree 3). The fact that the mass measure is diffuse and supported by the set of leaves makes the Brownian tree a *continuum random tree* according to Aldous' terminology (see Aldous [4]). For more detail on (1.13), see for instance Duquesne & Le Gall [12].

The choice of the normalizing constant  $\sqrt{2}$  for the underlying Brownian motion  $X$  is motivated by the following fact: let  $T_n^*$  be uniformly distributed on the set of rooted planar trees with  $n$  vertices. We view  $T_n^*$  as a graph embedded in the clockwise oriented upper half-plane, whose edges are segments of unit length and whose root is at the origin. Let us consider a particle that explores  $T_n^*$  as follows: it starts at the root and then it moves continuously on the tree at unit speed from the left to the right, backtracking as little as possible. During this exploration the particle visits each edge exactly twice and its journey lasts  $2(n-1)$  units of time. For all  $t \in [0, 2(n-1)]$ , we denote by  $C_t^{(n)}$  the distance between the root and the position of the particle at time  $t$ . The process  $(C_t^{(n)})_{t \in [0, 2(n-1)]}$  is called the *contour process* of  $T_n^*$ . Following an idea of Dwass [14], we can check that the contour process  $(C_t^{(n)})_{t \in [0, 2(n-1)]}$  is distributed as the (linear interpolation of the) simple random walk starting from 0, conditioned to stay nonnegative on  $[0, 2(n-1)]$  and to have the value 0 at time  $2(n-1)$ . Using a variant of Donsker's invariance principle, the rescaled contour function  $(n^{-\frac{1}{2}} C_{2(n-1)t}^{(n)})_{t \in [0, 1]}$  converges in law towards  $e$  under  $\mathbf{N}_{\text{nr}}$ : see for instance Le Gall [19]. Thus,

$$n^{-\frac{1}{2}} D_n^* \xrightarrow{\text{(law)}} D \quad \text{under } \mathbf{N}_{\text{nr}},$$

where  $D_n^*$  stands for the diameter of  $T_n^*$  and  $D$  is the diameter of the Brownian tree given by (1.11).

**Remark 1.2.** In the first paragraph of the introduction, we have introduced the random tree  $T_n$ , which is uniformly distributed on the set of rooted labelled trees with  $n$  vertices. The law of  $T_n$  is therefore distinct from that of  $T_n^*$ , which is uniformly distributed on the set of rooted planar trees with  $n$  vertices. Aldous [4] has proved that the tree  $T_n$ , whose graph distance is rescaled by a factor  $n^{-\frac{1}{2}}$ , converges to the tree coded by  $\sqrt{2}e$  under  $\mathbf{N}_{\text{nr}}$ . Thus,

$$\Delta \stackrel{\text{(law)}}{=} \sqrt{2}D \quad \text{under } \mathbf{N}_{\text{nr}}. \quad (1.14)$$

See Remark 1.5 below.

In this article, we prove the following result that characterizes the joint law of the height and diameter of the Brownian tree.

**Theorem 1.3.** Recall from (1.8) the definition of  $\mathbf{N}_{\text{nr}}$  and recall from (1.11) the definitions of  $\Gamma$  and  $D$ . We set

$$\forall \lambda, y, z \in (0, \infty), \quad L_\lambda(y, z) := \frac{1}{2\sqrt{\pi}} \int_0^\infty e^{-\lambda r} r^{-\frac{3}{2}} \mathbf{N}_{\text{nr}}(r^{\frac{1}{2}} D > 2y; r^{\frac{1}{2}} \Gamma > z) dr. \quad (1.15)$$

Note that

$$\forall \lambda, y, z \in (0, \infty), \quad L_1(y, z) = \lambda^{-\frac{1}{2}} L_\lambda(\lambda^{-\frac{1}{2}} y, \lambda^{-\frac{1}{2}} z). \quad (1.16)$$

Then,

$$L_1(y, z) = \coth(y \vee z) - 1 - \frac{1}{4} \mathbf{1}_{\{z \leq 2y\}} \frac{\sinh(2q) - 2q}{\sinh^4(y)}, \quad (1.17)$$

where  $q = y \wedge (2y - z)$ . In particular, this implies that

$$\forall \lambda, z \in (0, \infty), \quad L_\lambda(0, z) = \sqrt{\lambda} \coth(z\sqrt{\lambda}) - \sqrt{\lambda} \quad (1.18)$$

and

$$\forall \lambda, y \in (0, \infty), \quad L_\lambda(y, 0) = \sqrt{\lambda} \coth(y\sqrt{\lambda}) - \sqrt{\lambda} - \sqrt{\lambda} \frac{\sinh(2y\sqrt{\lambda}) - 2y\sqrt{\lambda}}{4 \sinh^4(y\sqrt{\lambda})}. \quad (1.19)$$

**Corollary 1.4.** For all  $y, z \in (0, \infty)$ , we set

$$\rho = z \vee \frac{y}{2} \quad \text{and} \quad \delta = \left( \frac{2(y-z)}{y} \vee 0 \right) \wedge 1. \quad (1.20)$$

Then we have

$$\mathbf{N}_{\text{nr}}(D > y; \Gamma > z) = 2 \sum_{n \geq 1} (2n^2 \rho^2 - 1) e^{-n^2 \rho^2} + \quad (1.21)$$

$$\sum_{n \geq 2} \binom{n+1}{3} \left[ [(n+\delta)^2 y^2 - 2] e^{-\frac{1}{4}(n+\delta)^2 y^2} - [(n-\delta)^2 y^2 - 2] e^{-\frac{1}{4}(n-\delta)^2 y^2} + \delta y (n^3 y^3 - 6ny) e^{-\frac{1}{4}n^2 y^2} \right]$$

and

$$\mathbf{N}_{\text{nr}}(D \leq y; \Gamma \leq z) = \frac{4\pi^{5/2}}{\rho^3} \sum_{n \geq 1} n^2 e^{-n^2 \pi^2 / \rho^2} - \quad (1.22)$$

$$\begin{aligned} & \frac{32\pi^{3/2}}{3} \sum_{n \geq 1} n \sin(2\pi n \delta) \left( \frac{2}{y^5} (2a_{n,y}^2 - 9a_{n,y} + 6) - \frac{3\delta^2 - 1}{y^3} (a_{n,y} - 1) \right) e^{-a_{n,y}} + \\ & \frac{16\pi^{1/2}}{3} \sum_{n \geq 1} \delta \cos(2\pi n \delta) \left( \frac{1}{y^3} (6a_{n,y}^2 - 15a_{n,y} + 3) - \frac{\delta^2 - 1}{2y} a_{n,y} \right) e^{-a_{n,y}} + \\ & \frac{16\pi^{1/2}}{3} \sum_{n \geq 1} \delta \left( \frac{1}{y^3} (4a_{n,y}^3 - 24a_{n,y}^2 + 27a_{n,y} - 3) + \frac{1}{2y} (2a_{n,y}^2 - 3a_{n,y}) \right) e^{-a_{n,y}}, \end{aligned}$$

where we have set  $a_{n,y} = 4(\pi n/y)^2$  for all  $y \in (0, \infty)$  and for all  $n \geq 1$  to simplify notation, and where  $\binom{n}{k} = \frac{n!}{k!(n-k)!}$  for all  $0 \leq k \leq n$ . In particular, (1.21) implies

$$\mathbf{N}_{\text{nr}}(\Gamma > y) = 2 \sum_{n \geq 1} (2n^2 y^2 - 1) e^{-n^2 y^2}, \quad (1.23)$$

and

$$\mathbf{N}_{\text{nr}}(D > y) = \sum_{n \geq 2} (n^2 - 1) \left( \frac{1}{6} n^4 y^4 - 2n^2 y^2 + 2 \right) e^{-n^2 y^2 / 4}. \quad (1.24)$$

On the other hand, (1.22) implies

$$\mathbf{N}_{\text{nr}}(\Gamma \leq y) = \frac{4\pi^{5/2}}{y^3} \sum_{n \geq 1} n^2 e^{-n^2 \pi^2 / y^2}, \quad (1.25)$$

and

$$\mathbf{N}_{\text{nr}}(D \leq y) = \frac{\sqrt{\pi}}{3} \sum_{n \geq 1} \left( \frac{8}{y^3} (24a_{n,y} - 36a_{n,y}^2 + 8a_{n,y}^3) + \frac{16}{y} a_{n,y}^2 \right) e^{-a_{n,y}}. \quad (1.26)$$

Thus the law of  $D$  under  $\mathbf{N}_{\text{nr}}$  has the following density:

$$f_D(y) = \frac{1}{12} \sum_{n \geq 1} (n^8 y^5 - n^6 y^3 (20 + y^2) + 20n^4 y (3 + y^2) - 60n^2 y) e^{-n^2 y^2/4} \quad (1.27)$$

$$= \frac{2\sqrt{\pi}}{3} \sum_{n \geq 1} \left( \frac{16}{y^4} (4a_{n,y}^4 - 36a_{n,y}^3 + 75a_{n,y}^2 - 30a_{n,y}) + \frac{8}{y^2} (2a_{n,y}^3 - 5a_{n,y}^2) \right) e^{-a_{n,y}} \quad (1.28)$$

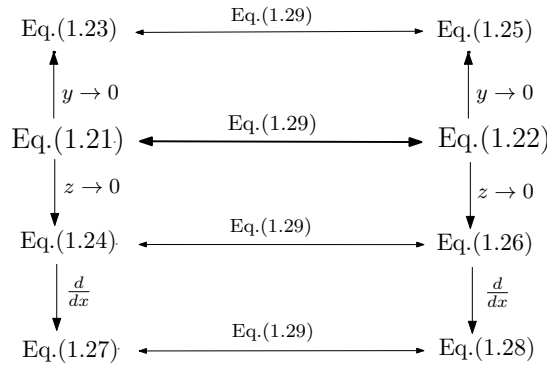


Figure 1: Relations between the equations in Corollary 1.4.

**Remark 1.5.** Let us briefly explain the logical connections between the eight assertions of Corollary 1.4, which are illustrated in Figure 1. We derive (1.22) from (1.21) using the following identity on the theta function due to Jacobi (1828), which is a consequence of the Poisson summation formula:

$$\forall t \in (0, \infty), \forall x, y \in \mathbb{C}, \quad \sum_{n \in \mathbb{Z}} e^{-(x+n)^2 t - 2\pi i n y} = e^{2\pi i x y} \left( \frac{\pi}{t} \right)^{\frac{1}{2}} \sum_{n \in \mathbb{Z}} e^{-\frac{\pi^2 (y+n)^2}{t} + 2\pi i n x}. \quad (1.29)$$

See for instance Weil [23], Chapter VII, Equation (12). Not surprisingly, (1.29) can also be used to derive (1.25) from (1.23), to derive (1.26) from (1.24), or to derive (1.28) from (1.27).

We obtain (1.27) (resp. (1.28)) by differentiating (1.24) (resp. (1.26)). By (1.14), we have

$$\forall y \in (0, \infty), \quad f_\Delta(y) = \frac{1}{\sqrt{2}} f_D \left( \frac{y}{\sqrt{2}} \right),$$

which immediately entails (1.2) from (1.28), since  $a_{n,y/\sqrt{2}} = 8(\pi n/y)^2 = b_{n,y}$ .

**Remark 1.6.** Recall that  $\Gamma = \max_{t \geq 0} e_t$ . Equations (1.23) and (1.25) are consistent with previous results on the distribution of the maximum of Brownian excursion: see for example Chung [10], though we need to keep in mind the difference between  $\mathbf{N}_{\text{nr}}$  and  $\mathbf{N}_{\text{Itô}}^+$ , as explained in Remark 1.1.

## 2 Preliminaries

**A geometric property on diameters of real trees.** We begin with a simple observation on the total height and diameter of a real tree.

**Lemma 2.1.** *Let  $(T, d, \rho)$  be a compact rooted real tree. Then  $\Gamma \leq D \leq 2\Gamma$ , where*

$$\Gamma = \sup_{u \in T} d(u, \rho) \quad \text{and} \quad D = \sup_{u, v \in T} d(u, v).$$

*Moreover, there exists a pair of points  $u_0, v_0 \in T$  with maximal distance. Namely,*

$$d(u_0, v_0) = \sup_{u, v \in T} d(u, v) = D. \quad (2.1)$$

*Then for any  $a \in T$ , we have*

$$\max \{d(u_0, a), d(v_0, a)\} = \sup_{u \in T} d(u, a). \quad (2.2)$$

*In particular,*

$$\max \{d(u_0, \rho), d(v_0, \rho)\} = \sup_{u \in T} d(u, \rho) = \Gamma. \quad (2.3)$$

*Proof.* The triangle inequality implies that  $D \leq 2\Gamma$  while the inequality  $\Gamma \leq D$  is a direct consequence of the definitions. As  $d : T^2 \rightarrow \mathbb{R}_+$  is continuous and  $T$  is compact, there exists a pair of points  $u_0, v_0 \in T$  such that (2.1) holds true. Note that (2.3) will follow from (2.2) by taking  $a$  to be the root  $\rho$ . So it remains to show (2.2). Let  $u, v \in T$ . Recall from the definition of real trees (given in the introduction) that  $\llbracket u, v \rrbracket$  stands for the unique geodesic path between  $u$  and  $v$ . To prove (2.2), let us assume without loss of generality that  $d(u_0, a) \geq d(v_0, a)$  and then show that  $d(u_0, a) \geq d(w, a)$  for any  $w \in T$ . Properties of real trees entail that there exists a point  $b \in T$  such that  $\llbracket a, u_0 \rrbracket \cap \llbracket a, v_0 \rrbracket = \llbracket a, b \rrbracket$  (see for instance Remark 2.21 and Lemma 2.22 in [15]). Similarly, there exists a point  $b' \in T$  such that  $\llbracket a, u_0 \rrbracket \cap \llbracket a, w \rrbracket = \llbracket a, b' \rrbracket$ . Here we enumerate the three possible locations of  $b'$  relative to  $b$ . See Figure 2.

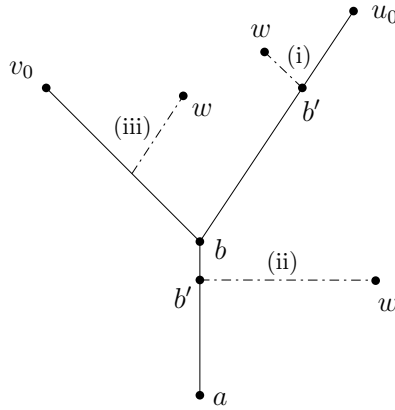


Figure 2: Three possibilities for  $(w, b')$

(i) Suppose that  $b' \in \llbracket b, u_0 \rrbracket$  and  $b' \neq b$ . Then we have  $d(b', u_0) \geq d(b', w)$ , since

$$d(v_0, u_0) = d(v_0, b') + d(b', u_0) \geq d(v_0, w) = d(v_0, b') + d(b', w)$$

by the definition of  $(u_0, v_0)$ . Thus,

$$d(a, u_0) = d(a, b') + d(b', u_0) \geq d(a, b') + d(b', w) = d(a, w).$$



- (ii) Suppose that  $b' \in \llbracket a, b \rrbracket$  and  $b' \neq b$ . Then we also have  $d(b', u_0) \geq d(b', w)$ . Indeed, since

$$d(v_0, u_0) = d(v_0, b) + d(b, u_0) \geq d(v_0, w) = d(v_0, b) + d(b, w),$$

it follows that  $d(b, u_0) \geq d(b, w)$ . Then,

$$d(b', u_0) > d(b, u_0) \geq d(b, w) > d(b', w).$$

By the same reason as before, we also get  $d(a, u_0) \geq d(a, w)$  in this case.

- (iii) Suppose that  $b' = b$ . Then we deduce from

$$d(u_0, v_0) = d(u_0, b) + d(b, v_0) \geq d(u_0, w) = d(u_0, b) + d(b, w)$$

that  $d(b, v_0) \geq d(b, w)$ . Combined with the assumption that  $d(a, u_0) \geq d(a, v_0)$ , this entails that

$$d(a, u_0) \geq d(a, v_0) = d(a, b) + d(b, v_0) \geq d(a, b) + d(b, w) = d(a, w).$$

In short, we have shown that  $d(a, u_0) \geq d(a, w)$  for any  $w \in \mathbb{T}$ . This completes the proof.  $\square$

**Williams' decomposition of Brownian excursions.** Let us recall the classical result of Williams' path decomposition of Brownian excursions (see for instance Revuz & Yor [21] Chapter XII Theorem 4.5). Define

$$\tau_* := \inf\{t > 0 : e_t = \max_{s \geq 0} e_s\}. \quad (2.4)$$

Then,

$$\mathbf{N}\text{-a.e. } (\mathbf{N}_{\text{nr}}\text{-a.s.}), \tau_* \text{ is the unique time at which } e \text{ reaches its maximum value.} \quad (2.5)$$

Recall from (1.11) the definition of the total height  $\Gamma$  of the Brownian tree coded by  $e$ . Then, we have  $\Gamma = e_{\tau_*}$ .

We also recall the distribution of  $\Gamma$  under  $\mathbf{N}$ :

$$\mathbf{N}(\Gamma \in dr) = \frac{dr}{r^2}. \quad (2.6)$$

See Revuz & Yor [21] Chapter XII Theorem 4.5 combined with Remark 1.1.

Williams's decomposition entails that there is a regular version of the family of conditioned laws  $\mathbf{N}(\cdot | \Gamma = r)$ ,  $r > 0$ . Namely,  $\mathbf{N}(\cdot | \Gamma = r)$ -a.s.  $\Gamma = r$ ,  $r \mapsto \mathbf{N}(\cdot | \Gamma = r)$  is weakly continuous on  $\mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$  and

$$\mathbf{N} = \int_0^\infty \mathbf{N}(\Gamma \in dr) \mathbf{N}(\cdot | \Gamma = r). \quad (2.7)$$

Let  $Z = (Z_t)_{t \geq 0}$  be a continuous process defined on the probability space  $(\Omega, \mathcal{F}, \mathbf{P})$  such that  $\frac{1}{\sqrt{2}}Z$  is distributed as a Bessel process of dimension 3 starting from 0. Let  $\tau_r = \inf\{t > 0 : Z_t = r\}$  be the hitting time of  $Z$  at level  $r \in (0, \infty)$ . We recall that

$$\forall \lambda \in \mathbb{R}_+, \quad \mathbf{E}[e^{-\lambda \tau_r}] = \frac{r\sqrt{\lambda}}{\sinh(r\sqrt{\lambda})}. \quad (2.8)$$

See Borodin & Salminen [8] Part II, Chapter 5, Section 2, Formula 2.0.1, p. 463, where we let  $x$  tend to 0 and take  $\alpha = \lambda$  and  $z = r/\sqrt{2}$ , since  $Z = \sqrt{2}R^{(3)}$ .

We next introduce the following notation

$$\overleftarrow{e}(t) = e_{(\tau_* - t)_+}; \quad \overrightarrow{e}(t) = e_{\tau_* + t}, \quad t \geq 0.$$

where  $(\cdot)_+$  stands for the positive part function. Williams' decomposition of Brownian excursion asserts that

for all  $r \in (0, \infty)$ , under  $\mathbf{N}(\cdot | \Gamma = r)$ , the two processes  $\overleftarrow{e}$  and  $\overrightarrow{e}$  are distributed as two independent copies of  $(Z_{(\tau_r - t)_+})_{t \geq 0}$ .

As a combined consequence of this decomposition and (2.8), we have

$$\forall r \in (0, \infty), \quad \mathbf{N}(e^{-\lambda \zeta} | \Gamma = r) = \mathbf{E}[e^{-\lambda \tau_r}]^2 = \left( \frac{r\sqrt{\lambda}}{\sinh(r\sqrt{\lambda})} \right)^2, \quad (2.9)$$

where we recall that  $\zeta$  stands for the lifetime of the excursion. Therefore,

$$\mathbf{N}(e^{-\lambda \zeta} \mathbf{1}_{\{\Gamma > a\}}) = \int_a^\infty \mathbf{N}(e^{-\lambda \zeta} | \Gamma = r) \mathbf{N}(\Gamma \in dr) = \int_a^\infty \frac{\lambda dr}{\sinh^2(r\sqrt{\lambda})} = \sqrt{\lambda} \coth(a\sqrt{\lambda}) - \sqrt{\lambda},$$

by (2.6) and (2.9). Recall from (1.5) that  $\mathbf{N}(1 - e^{-\lambda \zeta}) = \sqrt{\lambda}$ . Combining this fact with the above computation, we find that

$$\mathbf{N}(1 - e^{-\lambda \zeta} \mathbf{1}_{\{\Gamma \leq a\}}) = \sqrt{\lambda} \coth(a\sqrt{\lambda}). \quad (2.10)$$

This equation is used in the proof of Theorem 1.3.

**Spinal decomposition** Let us interpret Williams' decomposition in terms of a Poisson decomposition of the Brownian excursion. To that end, we need the following notation. Let  $h \in \mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$  have compact support. We assume that  $h(0) > 0$ . For all  $s \in \mathbb{R}_+$ , we set  $\underline{h}(s) = \inf_{0 \leq u \leq s} h(u)$ . Let  $(l_i, r_i)$ ,  $i \in \mathcal{I}(h)$  be the excursion intervals of  $h - \underline{h}$  away from 0; namely, they are the connected components of the open set  $\{s \geq 0 : h(s) - \underline{h}(s) > 0\}$ . For all  $i \in \mathcal{I}(h)$ , we next set

$$h^i(s) = (h - \underline{h})((l_i + s) \wedge r_i), \quad s \geq 0,$$

which is the excursion of  $h - \underline{h}$  corresponding to the interval  $(l_i, r_i)$ . Then we set

$$\mathcal{P}(h) = \sum_{i \in \mathcal{I}(h)} \delta_{(h(0) - h(l_i), h^i)},$$

that is a point measure on  $\mathbb{R}_+ \times \mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$ . We define

$$\mathcal{Q} := \mathcal{P}(\overleftarrow{e}) + \mathcal{P}(\overrightarrow{e}) =: \sum_{j \in \mathcal{J}} \delta_{(s_j, e^j)}. \quad (2.11)$$

We also introduce for all  $t \in (0, \infty)$  the following notation

$$\mathbf{N}_t = \mathbf{N}(\cdot \cap \{\Gamma \leq t\}). \quad (2.12)$$

The following lemma is a special case of a general result due to Abraham & Delmas [1].

**Lemma 2.2** (Proposition 1.1, Abraham & Delmas [1]). *Let  $r \in (0, \infty)$ . Then,  $\mathcal{Q}$  under  $\mathbf{N}(\cdot | \Gamma = r)$  is a Poisson point measure on  $\mathbb{R}_+ \times \mathbf{C}(\mathbb{R}_+, \mathbb{R}_+)$  with intensity measure  $2 \cdot \mathbf{1}_{[0, r]}(t) dt \mathbf{N}_t$ .*

**Interpretation in terms of the Brownian tree and consequences.** Let us interpret  $\mathcal{Q}$  in terms of a spinal decomposition of the Brownian tree  $\mathcal{T}$  coded by the Brownian excursion  $e$ . For other works using spinal decompositions, see for instance Aldous & Pitman [5], Haas, Pitman & Winkel [17] and the references therein. Recall that  $p : [0, \zeta] \rightarrow \mathcal{T}$  stands for the canonical projection and recall that  $\rho = p(0)$  is the root of  $\mathcal{T}$ . The point  $p(\tau_*)$  is the (unique) point of  $\mathcal{T}$  that attains the total height:  $d(\rho, p(\tau_*)) = \Gamma$ .

Denote by  $\mathcal{T}_{j'}^o$ ,  $j' \in \mathcal{J}'$ , the connected components of  $\mathcal{T} \setminus \llbracket \rho, p(\tau_*) \rrbracket$ . For all  $j' \in \mathcal{J}'$ , there exists a unique point  $\sigma_{j'} \in \llbracket \rho, p(\tau_*) \rrbracket$  such that  $\mathcal{T}_{j'} := \mathcal{T}_{j'}^o \cup \{\sigma_{j'}\}$  is the closure of  $\mathcal{T}_{j'}^o$  in  $\mathcal{T}$ . Recall the notation  $\mathcal{J}$  from (2.11). It is not difficult to see that  $\mathcal{J}'$  is in one-to-one correspondence with  $\mathcal{J}$ . Moreover, after a re-indexing, we can suppose that  $d(p(\tau_*), \sigma_j) = s_j$  and that  $(\mathcal{T}_j, d, \sigma_j)$  is the real tree coded by the excursion  $e^j$ , for each  $j \in \mathcal{J}$ . Then we set

$$\forall j \in \mathcal{J}, \quad \Gamma_j := \max_{s \geq 0} e^j(s) = \max_{\gamma \in \mathcal{T}_j} d(\sigma_j, \gamma), \quad (2.13)$$

that is the total height of the rooted real tree  $(\mathcal{T}_j, d, \sigma_j)$ . We claim that

$$\mathbf{N}\text{-a.e.} \quad D = \sup_{j \in \mathcal{J}} (s_j + \Gamma_j). \quad (2.14)$$

*Proof of (2.14).* First observe that for all  $t \in (0, \infty)$ ,  $\mathbf{N}_t$  is an infinite measure because  $\mathbf{N}$  is infinite and because  $\mathbf{N}(\Gamma > t) = 1/t$  by (2.6). By Lemma 2.2,  $\mathbf{N}$ -a.e. the closure of the set  $\{s_j; j \in \mathcal{J}\}$  is  $[0, \Gamma]$ . This entails that

$$\mathbf{N}\text{-a.e.} \quad \Gamma = \sup_{j \in \mathcal{J}} s_j \leq \sup_{j \in \mathcal{J}} (s_j + \Gamma_j). \quad (2.15)$$

Next, for all  $j \in \mathcal{J}$ , there exists  $\gamma_j \in \mathcal{T}_j$  such that  $d(\sigma_j, \gamma_j) = \Gamma_j$ . Then observe that

$$d(p(\tau_*), \gamma_j) = d(p(\tau_*), \sigma_j) + d(\sigma_j, \gamma_j) = s_j + \Gamma_j. \quad (2.16)$$

Note that Lemma 2.1 and (2.5) imply that  $D = \max_{\gamma \in \mathcal{T}} d(p(\tau_*), \gamma)$ . Comparing this with (2.16), we get

$$D \geq \sup_{j \in \mathcal{J}} (s_j + \Gamma_j). \quad (2.17)$$

On the other hand, there exists  $\gamma^* \in \mathcal{T}$  such that  $D = \max_{\gamma \in \mathcal{T}} d(p(\tau_*), \gamma) = d(p(\tau_*), \gamma^*)$  by Lemma 2.1. If  $\gamma^* \notin \llbracket \rho, p(\tau_*) \rrbracket$ , then there exists  $j^* \in \mathcal{J}$  such that  $\gamma^* \in \mathcal{T}_{j^*}$ . In consequence, we have  $D = d(p(\tau_*), \gamma^*) \leq s_{j^*} + \Gamma_{j^*}$ , and then  $D = \sup_{j \in \mathcal{J}} (s_j + \Gamma_j)$  when compared with (2.17). If  $\gamma^* \in \llbracket \rho, p(\tau_*) \rrbracket$ , then (2.15) implies that  $\gamma^* = \rho$  and  $D = \Gamma$ . In both cases (2.14) holds true.  $\square$

We next denote by  $\zeta_j$  the lifetime of  $e^j$  for all  $j \in \mathcal{J}$  and prove the following statement.

$$\mathbf{N}\text{-a.e.} \quad \sum_{j \in \mathcal{J}} \zeta_j = \zeta. \quad (2.18)$$

*Proof of (2.18).* Let  $\sigma \in \llbracket p(\tau_*), \rho \rrbracket$  be distinct from  $p(\tau_*)$  and  $\rho$ . Then  $n(\sigma) \geq 2$  and  $\sigma$  is not a leaf of  $\mathcal{T}$ . Recall from (1.12) the definition of the mass measure  $m$  and recall from (1.13) that  $\mathbf{N}_{\text{nr}}$ -a.s.  $m$  is diffuse and supported on the set of leaves of  $\mathcal{T}$ . By (1.7), this property also holds true  $\mathbf{N}$ -almost everywhere and we thus get

$$\mathbf{N}\text{-a.e.} \quad m(\llbracket p(\tau_*), \rho \rrbracket) = 0.$$

Recall that  $\mathcal{T}_j^o$ ,  $j \in \mathcal{J}$ , are the connected components of  $\mathcal{T} \setminus \llbracket \rho, p(\tau_*) \rrbracket$ . Thus,

$$\mathbf{N}\text{-a.e.} \quad m(\mathcal{T}) = m(\llbracket p(\tau_*), \rho \rrbracket) + \sum_{j \in \mathcal{J}} m(\mathcal{T}_j^o) = \sum_{j \in \mathcal{J}} m(\mathcal{T}_j^o). \quad (2.19)$$

Recall that  $\mathcal{T}_j = \mathcal{T}_j^o \cup \{\sigma_j\}$  and that  $m$  is  $\mathbf{N}$ -a.e. diffuse, which entails  $m(\mathcal{T}_j) = m(\mathcal{T}_j^o)$ , for all  $j \in \mathcal{J}$ . Moreover, since  $(\mathcal{T}_j, d, \sigma_j)$  is coded by the excursion  $e^j$ , we have  $\zeta_j = m(\mathcal{T}_j)$ . For a similar reason, we also have  $\zeta = m(\mathcal{T})$ . This, combined with (2.19), entails (2.18).  $\square$

### 3 Proof of Theorem 1.3

First we note that by (1.9),

$$L_\lambda(y, z) = \frac{1}{2\sqrt{\pi}} \int_0^\infty dr e^{-\lambda r} r^{-\frac{3}{2}} \mathbf{N}_{\text{nr}}(r^{\frac{1}{2}} D > 2y; r^{\frac{1}{2}} \Gamma > z) = \mathbf{N}(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y; \Gamma > z\}}). \quad (3.1)$$

Observe that the scaling property (1.16) is a direct consequence of the scaling property of  $\mathbf{N}$  (see (1.6)).

We next compute the right hand side of (3.1). To that end, recall from (2.11) the spinal decomposition of the excursion  $e$  and recall from (2.13) the notation  $\Gamma_j = \max_{s \geq 0} e^j(s)$ , for all  $j \in \mathcal{J}$ ; also recall that  $\zeta_j$  stands for the lifetime of  $e^j$ . Let  $r, y \in (0, \infty)$  be such that  $y \leq r \leq 2y$ . We apply successively (2.14), (2.18), Lemma 2.2 and Campbell's formula for Poisson point measures and find that

$$\begin{aligned} \mathbf{N}(e^{-\lambda\zeta} \mathbf{1}_{\{D \leq 2y\}} | \Gamma = r) &= \mathbf{N}\left(\prod_{j \in \mathcal{J}} e^{-\lambda\zeta_j} \mathbf{1}_{\{s_j + \Gamma_j \leq 2y\}} | \Gamma = r\right) \\ &= \exp\left(-2 \int_0^r dt \mathbf{N}_t(1 - e^{-\lambda\zeta} \mathbf{1}_{\{\Gamma \leq 2y-t\}})\right). \end{aligned} \quad (3.2)$$

Recall from (2.12) that  $\mathbf{N}_t = \mathbf{N}(\cdot \cap \{\Gamma \leq t\})$  and observe that

$$\int_0^r dt \mathbf{N}_t(1 - e^{-\lambda\zeta} \mathbf{1}_{\{\Gamma \leq 2y-t\}}) = \int_0^y dt \mathbf{N}((1 - e^{-\lambda\zeta}) \mathbf{1}_{\{\Gamma \leq t\}}) + \int_y^r dt \mathbf{N}(\mathbf{1}_{\{\Gamma \leq t\}} - e^{-\lambda\zeta} \mathbf{1}_{\{\Gamma < 2y-t\}}). \quad (3.3)$$

By (2.10) and by (2.6),

$$\mathbf{N}((1 - e^{-\lambda\zeta}) \mathbf{1}_{\{\Gamma \leq t\}}) = \mathbf{N}(1 - e^{-\lambda\zeta} \mathbf{1}_{\{\Gamma \leq t\}}) - \mathbf{N}(\Gamma > t) = \sqrt{\lambda} \coth(t\sqrt{\lambda}) - \frac{1}{t} \quad (3.4)$$

and

$$\begin{aligned} \mathbf{N}(\mathbf{1}_{\{\Gamma \leq t\}} - e^{-\lambda\zeta} \mathbf{1}_{\{\Gamma < 2y-t\}}) &= \mathbf{N}(1 - e^{-\lambda\zeta} \mathbf{1}_{\{\Gamma \leq 2y-t\}}) - \mathbf{N}(\Gamma > t) \\ &= \sqrt{\lambda} \coth((2y-t)\sqrt{\lambda}) - \frac{1}{t}. \end{aligned} \quad (3.5)$$

Then observe that for all  $\varepsilon, a \in (0, \infty)$  such that  $\varepsilon < a$ ,

$$\int_\varepsilon^a \left( \sqrt{\lambda} \coth(t\sqrt{\lambda}) - \frac{1}{t} \right) dt = \log\left(\frac{\sinh(a\sqrt{\lambda})}{a}\right) - \log\left(\frac{\sinh(\varepsilon\sqrt{\lambda})}{\varepsilon}\right).$$

Thus, as  $\varepsilon \rightarrow 0$ , we get

$$\forall a \in \mathbb{R}_+, \quad \int_0^a \left( \sqrt{\lambda} \coth(t\sqrt{\lambda}) - \frac{1}{t} \right) dt = \log\left(\frac{\sinh(a\sqrt{\lambda})}{a\sqrt{\lambda}}\right). \quad (3.6)$$

A direct computation based on (3.6), combined with (3.3), (3.4), (3.5) and (3.2), entails

$$\mathbf{N}(e^{-\lambda\zeta} \mathbf{1}_{\{D \leq 2y\}} | \Gamma = r) = \frac{(r\sqrt{\lambda} \sinh((2y-r)\sqrt{\lambda}))^2}{\sinh^4(y\sqrt{\lambda})}.$$

Combining this with (2.9), we get for all  $r, y \in (0, \infty)$  satisfying  $y \leq r \leq 2y$ ,

$$\mathbf{N}(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y\}} | \Gamma = r) = \left( \frac{r\sqrt{\lambda}}{\sinh(r\sqrt{\lambda})} \right)^2 - \frac{(r\sqrt{\lambda} \sinh((2y-r)\sqrt{\lambda}))^2}{\sinh^4(y\sqrt{\lambda})}. \quad (3.7)$$

Next, let  $r, y \in (0, \infty)$  be such that  $r > 2y$ . By Lemma 2.1,  $\Gamma \leq D \leq 2\Gamma$ . Therefore,

$$\forall r, y \in (0, \infty) : r > 2y, \quad \mathbf{N}\left(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y\}} \mid \Gamma = r\right) = \mathbf{N}\left(e^{-\lambda\zeta} \mid \Gamma = r\right) = \left(\frac{r\sqrt{\lambda}}{\sinh(r\sqrt{\lambda})}\right)^2. \quad (3.8)$$

Finally, let  $r < y$ . Then  $\mathbf{N}(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y\}} \mid \Gamma = r) = 0$ , since  $\Gamma \leq D \leq 2\Gamma$ . Combining this with (3.7) and (3.8), we obtain that

$$\begin{aligned} \mathbf{N}\left(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y, \Gamma > z\}}\right) &= \int_z^\infty \mathbf{N}\left(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y\}} \mid \Gamma = r\right) \mathbf{N}(\Gamma \in dr) \\ &= \int_{z \vee y}^{2y \vee z} \mathbf{N}\left(e^{-\lambda\zeta} \mathbf{1}_{\{D > 2y\}} \mid \Gamma = r\right) \mathbf{N}(\Gamma \in dr) + \int_{2y \vee z}^\infty \mathbf{N}\left(e^{-\lambda\zeta} \mid \Gamma = r\right) \mathbf{N}(\Gamma \in dr) \\ &= \sqrt{\lambda} \left( \coth((z \vee y)\sqrt{\lambda}) - 1 \right) - \mathbf{1}_{\{z \leq 2y\}} \frac{\sqrt{\lambda} \sinh(2q\sqrt{\lambda}) - 2\lambda q}{4 \sinh^4(y\sqrt{\lambda})}, \end{aligned}$$

where we recall the notation  $q = y \wedge (2y - z)$ . By (3.1), this concludes the proof of Theorem 1.3.

#### 4 Proof of Corollary 1.4

We introduce the following notation for the Laplace transform on  $\mathbb{R}_+$ : for all the Lebesgue integrable functions  $f : \mathbb{R}_+ \rightarrow \mathbb{R}$ , we set

$$\forall \lambda \in \mathbb{R}_+, \quad \mathcal{L}_\lambda(f) := \int_0^\infty dx e^{-\lambda x} f(x),$$

which is well-defined. Note that if  $f, g$  are two continuous and integrable functions such that  $\mathcal{L}_\lambda(f) = \mathcal{L}_\lambda(g)$  for all  $\lambda \in [0, \infty)$ , then we have  $f = g$ , by the injectivity of the Laplace transform and standard arguments.

For all  $a, x \in (0, \infty)$ , we set  $f_a(x) = \frac{a}{2\sqrt{\pi}} x^{-3/2} e^{-a^2/4x}$ . It is well-known that  $\mathcal{L}_\lambda(f_a) = e^{-a\sqrt{\lambda}}$  for all  $\lambda \in \mathbb{R}_+$  (see for instance Borodin & Salminen [8] Appendix 3, Particular formulæ 2, p. 650). Then we set

$$\begin{aligned} g_a(x) &= \frac{\partial}{\partial x} f_a(x) = \frac{1}{8\sqrt{\pi}} x^{-7/2} e^{-\frac{a^2}{4x}} (a^3 - 6ax) \quad \text{and} \\ h_a(x) &= -\frac{\partial}{\partial a} f_a(x) = \frac{1}{4\sqrt{\pi}} x^{-5/2} e^{-\frac{a^2}{4x}} (a^2 - 2x). \end{aligned}$$

Consequently, for all  $\lambda \in \mathbb{R}_+$ ,

$$\mathcal{L}_\lambda(g_a) = \lambda e^{-a\sqrt{\lambda}} \quad \text{and} \quad \mathcal{L}_\lambda(h_a) = \sqrt{\lambda} e^{-a\sqrt{\lambda}}. \quad (4.1)$$

(See also Borodin & Salminen [8] Appendix 3, Particular formulæ 3 and 4, p. 650.) Moreover, we have the following bounds: for all  $\lambda \in \mathbb{R}_+$ ,

$$\mathcal{L}_\lambda(|g_a|) \leq \frac{1}{8\sqrt{\pi}} \int_0^\infty dx e^{-\lambda x} x^{-7/2} e^{-\frac{a^2}{4x}} (a^3 + 6ax) = \lambda e^{-a\sqrt{\lambda}} + \frac{6}{a} \sqrt{\lambda} e^{-a\sqrt{\lambda}} + \frac{6}{a^2} e^{-a\sqrt{\lambda}}, \quad (4.2)$$

$$\mathcal{L}_\lambda(|h_a|) \leq \frac{1}{4\sqrt{\pi}} \int_0^\infty dx e^{-\lambda x} x^{-5/2} e^{-\frac{a^2}{4x}} (a^2 + 2x) = \sqrt{\lambda} e^{-a\sqrt{\lambda}} + \frac{2}{a} e^{-a\sqrt{\lambda}}. \quad (4.3)$$

Let  $y, z \in (0, \infty)$ . Recall from (1.20) the notation  $\rho$  and  $\delta$ . Note that  $(1 - x)^{-4} =$

$\sum_{n \geq 0} \binom{n+3}{3} x^n$ , for all  $x \in [0, 1)$ . Then (1.17) implies that

$$\begin{aligned} L_1(\tfrac{1}{2}y, z) &= \coth \rho - 1 - \frac{\sinh(\delta y) - \delta y}{4 \sinh^4(y/2)} = \frac{2e^{-2\rho}}{1 - e^{-2\rho}} + \frac{2e^{-2y}(e^{-\delta y} - e^{\delta y})}{(1 - e^{-y})^4} + \frac{4\delta y e^{-2y}}{(1 - e^{-y})^4} \\ &= \sum_{n \geq 1} 2e^{-2n\rho} + \sum_{n \geq 0} 2 \binom{n+3}{3} \left( e^{-(n+2+\delta)y} - e^{-(n+2-\delta)y} + 2\delta y e^{-(n+2)y} \right) \\ &= \sum_{n \geq 1} 2e^{-2n\rho} + \sum_{n \geq 2} 2 \binom{n+1}{3} \left( e^{-(n+\delta)y} - e^{-(n-\delta)y} + 2\delta y e^{-ny} \right). \end{aligned}$$

Thus, by (1.16), we obtain that

$$\begin{aligned} \frac{1}{2\sqrt{\pi}} \int_0^\infty e^{-\lambda r} r^{-\frac{3}{2}} \mathbf{N}_{\text{nr}}(r^{\frac{1}{2}}D > y; r^{\frac{1}{2}}\Gamma > z) dr &= L_\lambda(\tfrac{1}{2}y, z) = \sqrt{\lambda} L_1(\tfrac{1}{2}y\sqrt{\lambda}, z\sqrt{\lambda}) \\ &= \sum_{n \geq 1} 2\sqrt{\lambda} e^{-2n\rho\sqrt{\lambda}} + \sum_{n \geq 2} 2 \binom{n+1}{3} \left( \sqrt{\lambda} e^{-(n+\delta)y\sqrt{\lambda}} - \sqrt{\lambda} e^{-(n-\delta)y\sqrt{\lambda}} + 2\delta y \lambda e^{-ny\sqrt{\lambda}} \right) \\ &= \sum_{n \geq 1} 2\mathcal{L}_\lambda(h_{2n\rho}) + \sum_{n \geq 2} 2 \binom{n+1}{3} \mathcal{L}_\lambda(h_{(n+\delta)y} - h_{(n-\delta)y} + 2\delta y g_{ny}). \end{aligned} \quad (4.4)$$

Observe that for all  $r \in \mathbb{R}_+$ ,

$$\sum_{n \geq 1} 2 \sup_{[0, r]} |h_{2n\rho}| + \sum_{n \geq 2} 2 \binom{n+1}{3} \left( \sup_{[0, r]} |h_{(n+\delta)y}| + \sup_{[0, r]} |h_{(n-\delta)y}| + 2\delta y \sup_{[0, r]} |g_{ny}| \right) < \infty. \quad (4.5)$$

Then, for any  $r \in \mathbb{R}_+$ , we set

$$\phi_{y,z}(r) = 2 \sum_{n=1}^\infty h_{2n\rho}(r) e^{-r} + \sum_{n=2}^\infty 2 \binom{n+1}{3} \left( h_{(n+\delta)y}(r) e^{-r} - h_{(n-\delta)y}(r) e^{-r} + 2\delta y g_{ny}(r) e^{-r} \right),$$

which is well-defined and continuous thanks to (4.5). The bounds (4.2) and (4.3) imply that  $\phi_{y,z}$  is Lebesgue integrable. Moreover, (4.4) asserts that  $L_{\lambda+1}(\tfrac{1}{2}y, z) = \mathcal{L}_\lambda(\phi_{y,z})$ . By the injectivity of the Laplace transform for continuous integrable functions (as mentioned above), we get

$$\forall r \in \mathbb{R}_+, \quad \phi_{y,z}(r) = \frac{1}{2\sqrt{\pi}} e^{-r} r^{-\frac{3}{2}} \mathbf{N}_{\text{nr}}(r^{\frac{1}{2}}D > y; r^{\frac{1}{2}}\Gamma > z),$$

which entails (1.21) by taking  $r = 1$ .

Since  $\Gamma \leq D \leq 2\Gamma$ , if  $z = y$ , then  $\mathbf{N}_{\text{nr}}(D > y; \Gamma > y) = \mathbf{N}_{\text{nr}}(\Gamma > y)$  and (1.21) immediately implies (1.23) because in this case  $\rho = y$  and  $\delta = 0$ . If  $z = y/2$ , then  $\mathbf{N}_{\text{nr}}(D > y; \Gamma > y/2) = \mathbf{N}_{\text{nr}}(D > y)$ ,  $\rho = y/2$ ,  $\delta = 1$  and (1.21) implies

$$\begin{aligned} \mathbf{N}_{\text{nr}}(D > y) &= \sum_{n \geq 1} (n^2 y^2 - 2) e^{-\frac{1}{4}n^2 y^2} + \sum_{n \geq 2} \binom{n+1}{3} \left[ [(n+1)^2 y^2 - 2] e^{-\frac{1}{4}(n+1)^2 y^2} - \right. \\ &\quad \left. [(n-1)^2 y^2 - 2] e^{-\frac{1}{4}(n-1)^2 y^2} + y(n^3 y^3 - 6ny) e^{-\frac{1}{4}n^2 y^2} \right], \end{aligned}$$

which entails (1.24) by re-indexing the sums according to  $e^{-n^2 y^2/4}$ ; we leave the detail to the reader. We next derive (1.27) by differentiating (1.24). As mentioned in Remark 1.5, we use Jacobi identity (1.29) to derive (1.22) from (1.21). The computations are long but straightforward: we leave them to the reader. Finally, for the same reason as before, (1.22) entails (1.25) by taking  $\rho = y$  and  $\delta = 0$ . It also entails (1.26) by taking  $\rho = \frac{1}{2}y$  and  $\delta = 1$ . Differentiating (1.26) gives (1.28). This completes the proof of Corollary 1.4.

## References

- [1] Romain Abraham and Jean-François Delmas, *Williams' decomposition of the Lévy continuum random tree and simultaneous extinction probability for populations with neutral mutations*, Stochastic Process. Appl. **119** (2009), no. 4, 1124–1143. MR-2508567
- [2] David Aldous, *The continuum random tree. I*, Ann. Probab. **19** (1991), no. 1, 1–28. MR-1085326
- [3] ———, *The continuum random tree. II. An overview*, Stochastic analysis (Durham, 1990), London Math. Soc. Lecture Note Ser., vol. 167, Cambridge Univ. Press, Cambridge, 1991, pp. 23–70. MR-1166406
- [4] ———, *The continuum random tree. III*, Ann. Probab. **21** (1993), no. 1, 248–289. MR-1207226
- [5] David Aldous and Jim Pitman, *Tree-valued Markov chains derived from Galton-Watson processes*, Ann. Inst. H. Poincaré Probab. Statist. **34** (1998), no. 5, 637–686. MR-1641670
- [6] Jean Bertoin, *Lévy processes*, Cambridge Tracts in Mathematics, vol. 121, Cambridge University Press, Cambridge, 1996. MR-1406564
- [7] Robert M. Blumenthal, *Excursions of Markov processes*, Probability and its Applications, Birkhäuser Boston, Inc., Boston, MA, 1992. MR-1138461
- [8] Andrei N. Borodin and Paavo Salminen, *Handbook of Brownian motion—facts and formulae*, Probability and its Applications, Birkhäuser Verlag, Basel, 1996. MR-1477407
- [9] Nicolas Broutin and Philippe Flajolet, *The distribution of height and diameter in random non-plane binary trees*, Random Structures Algorithms **41** (2012), no. 2, 215–252. MR-2956055
- [10] Kai Lai Chung, *Excursions in Brownian motion*, Ark. Mat. **14** (1976), no. 2, 155–177. MR-0467948
- [11] Andreas Dress, Vincent Moulton, and Werner Terhalle, *T-theory: an overview*, European J. Combin. **17** (1996), no. 2-3, 161–175, Discrete metric spaces (Bielefeld, 1994). MR-1379369
- [12] Thomas Duquesne and Jean-François Le Gall, *Probabilistic and fractal aspects of Lévy trees*, Probab. Theory Related Fields **131** (2005), no. 4, 553–603. MR-2147221
- [13] Thomas Duquesne and Minmin Wang, *Decomposition of Lévy trees along their diameter*, to appear in Ann. Inst. H. Poincaré Probab. Statist. <http://arxiv.org/abs/1503.05069>, 2015.
- [14] Meyer Dwass, *Branching processes in simple random walk*, Proc. Amer. Math. Soc. **51** (1975), 270–274. MR-0370775
- [15] Steven N. Evans, *Probability and real trees*, Lecture Notes in Mathematics, vol. 1920, Springer, Berlin, 2008, Lectures from the 35th Summer School on Probability Theory held in Saint-Flour, July 6–23, 2005. MR-2351587
- [16] Steven N. Evans, Jim Pitman, and Anita Winter, *Rayleigh processes, real trees, and root growth with re-grafting*, Probab. Theory Related Fields **134** (2006), no. 1, 81–126. MR-2221786
- [17] Bénédicte Haas, Jim Pitman, and Matthias Winkel, *Spinal partitions and invariance under re-rooting of continuum random trees*, Ann. Probab. **37** (2009), no. 4, 1381–1411. MR-2546748
- [18] Jean-François Le Gall, *The uniform random tree in a Brownian excursion*, Probab. Theory Related Fields **96** (1993), no. 3, 369–383. MR-1231930
- [19] ———, *Random trees and applications*, Probab. Surv. **2** (2005), 245–311. MR-2203728
- [20] J. Pitman, *Combinatorial stochastic processes*, Lecture Notes in Mathematics, vol. 1875, Springer-Verlag, Berlin, 2006, Lectures from the 32nd Summer School on Probability Theory held in Saint-Flour, July 7–24, 2002, With a foreword by Jean Picard. MR-2245368
- [21] Daniel Revuz and Marc Yor, *Continuous martingales and Brownian motion*, third ed., Grundlehren der Mathematischen Wissenschaften [Fundamental Principles of Mathematical Sciences], vol. 293, Springer-Verlag, Berlin, 1999. MR-1725357
- [22] G. Szekeres, *Distribution of labelled trees by diameter*, Combinatorial mathematics, X (Adelaide, 1982), Lecture Notes in Math., vol. 1036, Springer, Berlin, 1983, pp. 392–397. MR-731595
- [23] André Weil, *Elliptic functions according to Eisenstein and Kronecker*, Springer-Verlag, Berlin-New York, 1976, Ergebnisse der Mathematik und ihrer Grenzgebiete, Band 88. MR-0562289

**Acknowledgments.** This work is partly supported by the grant ANR-14-CE25- 0014 (ANR GRAAL). I am deeply grateful to Thomas Duquesne for suggesting this problem and for those fruitful discussions. I would like to thank Romain Abraham and Svante Janson for a careful reading of an earlier version. I am also thankful to an anonymous referee whose suggestions have helped improve the writing.



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